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This is the author's final, accepted and refereed manuscript to the article published in

***Journal of Urban Economics*, 69(2011)3: 319-327**

DOI: <http://dx.doi.org/10.1016/j.jue.2010.12.002>

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Evaluating Efficient Public Good Provision: Theory and Evidence from a Generalised Conditional Efficiency Model for Public Libraries*

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November 30, 2010

Abstract

Provision of most public goods (e.g., health care, libraries, education, police, fire protection, utilities) can be characterised by a two-stage production process. In the first stage, basic inputs (e.g., labour and capital) are used to generate service potential (e.g., opening hours, materials), which is then, in the second stage, transformed into observed outputs (e.g., school outcomes, library circulation, crimes solved). As final outputs are also affected by demand-side factors, conflating both production stages likely leads to biased inferences about public productive (in)efficiency and its determinants. Hence, this paper uses a specially tailored, fully non-parametric efficiency model allowing for both outlying observations and heterogeneity to analyse efficient public good provision in stage one only. We thereby employ a dataset comprising all 290 Flemish public libraries. Our findings indicate a statistically significant link between the ideological stance of the local government, wealth and density of the local population and source of library funding (i.e., local funding versus intergovernmental transfers) and library productive efficiency (though the casual nature of this correlation cannot be fully ascertained at present).

Keywords: Public good provision, Conditional efficiency, Nonparametric estimation, Libraries, Local government.

*The authors are grateful to the editor, two anonymous referees, Jos Blank, Elbert Dijkgraaf, Beate Jochimsen, Mika Kortelainen, Erwin Ooghe and seminar participants at WZB Berlin for useful comments and suggestions. The usual caveat applies.

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1 Introduction

In much the same way that concerns over allocative efficiency are at the heart of micro-economic theory (e.g., Leibenstein, 1966; Frantz, 1992), allocative efficiency in the public sector has always been a major concern in public finance. Numerous studies, for example, analyze whether local governments - which often have important responsibilities with respect to education, housing, health care, social welfare, recreation, infrastructure and the environment (including refuse collection) (John, 2001) - have a tendency to over- or underprovide public goods (see, e.g., the pioneering work of Brueckner, 1979, 1982, 1983 and many references thereto). Moreover, scholars studying the decentralization of tasks from higher-level governments to the local level often evaluate this evolution in terms of allocative efficiency. Smaller jurisdictions with more homogeneous populations are argued to increase allocative efficiency as they are more capable of matching the provision of public goods with the preferences of their constituents (Musgrave, 1959; Oates, 1972), while numerous "*informal and formal versions of the Tiebout model demonstrate that private allocative efficiency tends to be increased by Tiebout choice*" (Hoxby, 2000, 1211).

In contrast, this paper concentrates on local government productive efficiency.¹ This has received significantly less attention in the decentralization literature thus far (for important exceptions, see, e.g., Hoxby, 1999, 2000), even though one could argue that decentralization is most fruitful when local governments are, all else equal, more productively efficient than higher-level governments (e.g., Geys and Moesen, 2009). This relative neglect is all the more surprising given that the financial constraints within which local governments are expected to execute their (increasing) assignments have tightened significantly over the past decades. Indeed, given that tax- and deficit-increases are often politically costly (e.g., Geys and Vermeir, 2008a, b), one way to deal with increasing tasks *and* tightening budget requirements is to improve productive or technical efficiency (understood in terms of providing a maximum amount of output for a given level of inputs; see Koopmans, 1951; Fried *et al.*, 2008).

We are clearly not the first attempting to measure and explain local government productive efficiency (for reviews, see Tang, 1997; De Borger and Kerstens, 2000). Yet, we differ from this previous body of work in three crucial respects. First, we build on important - but often neglected - insights from the urban governance and public administration literatures to more thoroughly describe the public sector production process prior to the actual analysis. These

¹As in the private sector, efficiency constitutes one among many aims; including effectiveness, equity, responsiveness, adequateness and appropriateness (Dunn, 2004). Our focus on productive efficiency obviously does not imply that it should take precedence over other aims of public service provision. Note also that we will use the terms productive and technical efficiency interchangeably throughout the paper.

literatures illustrate that effective public service provision depends on an active involvement by the recipient of these services (e.g., Whitaker, 1980; Parks et al., 1981; Kiser, 1984; Parry, 1996). For example, schools can "*supply little education without inputs from students*", while police forces have "*very little capacity to affect community safety and security without citizen input*" such as reporting crimes or testifying in court (Parks et al., 1981, 1003). Such ‘coproduction’ has important implications for the measurement of technical efficiency, as it suggests that observable outcomes (e.g., library circulation, school results, waste collected, fires extinguished, crimes solved) - the most commonly employed output indicator in existing studies of public sector productive efficiency² - are inappropriate as they are not really ‘produced’ in a strict sense by the public service provider (see also Cordero-Ferrera et al., 2008). We therefore propose to view public good provision as a two-stage production process (adapted from Hammond, 2002) in which, *first*, basic inputs – such as labour and capital – are translated into ‘service potential’ – such as available materials and opening hours – and then, *secondly*, the latter are transformed into observable outputs – such as school outcomes, library circulation or crimes solved. Particularly in the first stage of this process can the public service producer be most directly held accountable for translating a given amount of public expenditures into a maximum possible amount of service potential (whereas the second stage is probably more appropriately analyzed in a supply-demand framework).

As a second contribution, we employ a recently developed fully non-parametric framework and thus do not impose any *a priori* assumption on the production technology. This is crucial given the difficulty - if not impossibility - to argue that the public good production process follows one or another functional form. While our approach is closely related to Data Envelopment Analysis (DEA) models (Charnes *et al.*, 1978; Deprins *et al.*, 1984), it goes further than such models by allowing for outliers (following the order-*m* technique of Cazals *et al.*, 2002) and heterogeneity (building on the conditional efficiency estimators of Daraio and Simar, 2005, 2007). Note that reliance on such conditional efficiency estimates is particularly convenient as it does not require a separability condition (i.e., the assumption that the exogenous environment does not influence the level of basic inputs and service potential). The final model is based on De Witte and Kortelainen (2008), who extended Daraio and Simar (2005, 2007) to allow for (1) both discrete and continuous exogenous variables and (2) statistical inference in the conditional efficiency approach. As such, besides reducing the impact of outliers and controlling for heterogeneity, we are able to non-parametrically evaluate the strength (though not necessarily the causal nature) of the correlation between exogenous characteristics and productive efficiency.

The latter also constitutes our third contribution. Previous studies generally fail to evaluate how the institutional environment – in terms of socio-demographic, economic or political

²See, for example, Worthington (2001) for the case of education.

characteristics – affects efficiency, or look at this via an econometric two-stage approach (e.g., De Borger *et al.*, 1994; Geys, 2006; Hemmeter, 2006; Borge *et al.*, 2008). Both exclusion of such background factors and their use in a two-stage approach, however, leads to biased results and incorrect inferences (see, respectively, Battese and Coelli, 1995 and Kumbhakar *et al.*, 1991; Reifschneider and Stevenson, 1991). In this paper, we exploit the above-mentioned non-parametric conditional efficiency model to include the operational environment immediately in the efficiency estimates. Relying on extensive public choice and political economics literatures, we thereby focus on the following elements: (1) ideological stance of the local government, (2) share of women in the local council, (3) wealth of the municipality, (4) population concentration, and (5) source of public funding.

While our central argument - and the ensuing empirical approach - can be readily applied to various public goods, our empirical application exploits an exceptionally rich dataset of (all 290) municipal public libraries in Flanders in 2007.³ The Flemish setting is particularly attractive since nearly every municipality has its own library, generating a large and diverse dataset. Moreover, as the central and regional governments in Belgium set the overall framework in which local public service providers operate, the latter’s work is largely execution-oriented and devoid of value choices (in contrast to, say, the US, where the value-component of local policy decisions is larger). This generates a situation that is particularly conducive to efficiency measurements as the value-content or neutrality of the inputs and outputs then becomes less of an issue (see also Geys and Moesen, 2009). Finally, we focus on libraries as local public library services are unlikely to be essential to individuals’ choice of residence (for recent evidence, see Bhatt, 2010), unlike, for example, a jurisdictions’ public education, tax policy or public safety. Moreover, selection of consumers by public libraries is unlikely to occur (unlike in, for example, education or health care; e.g., Parry, 1996). This is important since it strongly mitigates potential concerns about endogeneity and identification (more details below). Our findings suggest that the ideological stance of the local government, the wealth and density of the local population and the source of library funding (i.e., local funding versus intergovernmental transfers) are significantly correlated with an efficient generation of service potential. At odds with recent work on the effects of female representation on public policy (e.g., Chattopadhyay and Duflo, 2004; Geys and Revelli, 2009; Svaleryd, 2009), the number of women in the local government or the presence of a female mayor does not add to the explanatory power of the model, *ceteris paribus*. Still, although, as mentioned, our case selection mitigates potential concerns about endogeneity and identification, care should be taken to interpret these findings as causal relations (we return to this below).

³To facilitate the application in alternative settings, the R code underlying the present analysis is available from the authors upon request.

The remainder of the paper is structured as follows. Section 2 describes the theoretical background and our main hypotheses. Section 3 introduces the estimation methodology, while Section 4 discusses the institutional setting and data. Our findings are presented in Section 5. Finally, the main conclusions are summarized in Section 6.

2 Theoretical background and hypotheses

2.1 Public good provision as a two-stage production process

A key characteristic of many public services is that "*without the productive activities of consumers nothing of value will result*" (Parks et al., 1981, 1002). Such view of the importance of citizens-consumers as ‘co-producers’ of public service production and delivery first developed among urban governance and public administration scholars in the early 1980s (e.g., Whitaker, 1980; Parks et al., 1981; Kiser, 1984; Percy, 1984). Although discussion about the exact nature, origins and consequences of such coproduction continues (e.g., Alford, 2002; Mitlin, 2008), the basic idea is that effective public service provision depends on active involvement by the recipient of these services. That is, schools require students’ (and parents’) effort to produce decent exam results, health care provision can only succeed if doctors’ and nurses’ orders are followed by patients, waste collection proceeds faster when citizens appropriately bag it and transport it to the curb, unemployment assistance programs stand or fall with the active engagement of the (long-term) unemployed, ... Urban scholars have extensively discussed similar relevance of citizen involvement for a wide and varied range of local public services including fire and police protection, libraries, tax collection, recreation, and so on (for a review, see Percy, 1984).

Citizens’ coproduction represents a convenient way of pointing out that public goods have the consumers in their production function. This, however, has important implications for measuring public service providers’ technical or productive efficiency. Indeed, active involvement by the recipient of the service implies that observable outcomes (e.g., library circulation, school results, crimes resolved, fires extinguished and so on) are not really ‘produced’ in a strict sense by the public service provider - and thus are inappropriate measures to evaluate their technical (in)efficiency (see also Cordero-Ferrera et al., 2008). Indeed, given the importance of demand-side factors in the service production process, relying on observable outcomes in productive efficiency analyses may lead to strongly biased inferences of efficiency (for empirical evidence, see De Witte and Geys, 2010). For example, when observed library circulation (i.e., the final outcome) is low, a relatively high-cost library will appear inefficient when using circulation as the output variable in the analysis. Yet, it may at the same time be very efficient in translating its basic inputs (such as labour and capital) into books, opening hours and so on. If so, using circulation as an output measure will lead it to be unduly

described as productively inefficient simply because it suffers from low demand in its area. This is *not* to say that one should support locating such high-cost libraries in low-demand areas. Clearly, this would be a waste of public resources (i.e. *allocative* inefficiency). Rather, the argument is that, from a purely *productive* efficiency perspective, this library should be recognized as being technically efficient - and not be described as an underperformer for an element beyond its control (i.e., public demand for its service).

To accurately evaluate public sector technical efficiency and its determinants, we argue that one should concentrate on that part of the production process that is fully under the control of the service provider.⁴ We therefore propose to view public good provision as characterized by a two-stage production process (adapted from Hammond, 2002). In a first stage, basic inputs – such as (expenditures on) labour and capital – are employed in the production of what could be described as ‘service potential’ (Bookstein, 1981; Hammond, 2002). For public libraries, one can think of, for example, collection size and opening hours; in education, it may include teaching hours, teaching materials, school library; in health care, one can think of the number of hospital beds and operating rooms, opening hours, machinery (and similarly for other types of public services such as police, fire protection, water services, waste collection, and so on). In the second stage, this service potential is then transformed into observable outputs. For libraries, this could reflect book circulation and request processing, for health services the number of patients cured, and for fire and police protection the amount of fires extinguished and crimes solved, respectively. At this stage, the outcome is clearly not solely determined by the public service provider, but also depends to an important extent on the ‘demand’ in a given area for the services provided. For example, fire services are only provided if there are fires to be put out, while good exam results require student input and library circulation a population that cares about reading. Evidently, our distinction between service potential and final outputs is reminiscent of the distinction proposed by Becker (1965) and Lancaster (1966) between goods and commodities; where goods are seen as intermediate outcomes provided by firms and subsequently transformed by consumers into commodities that fulfil their desires. In such view, food items in a supermarket can be seen as goods, while individuals’ meals or nutrition are the commodity (or final outcome) (see also Kiser, 1984).

This distinction between two production phases of public good provision allows analyses of technical efficiency to concentrate on that stage of public good production where the service provider has full control over the inputs and outputs (i.e., stage one). At this stage, technical efficiency can most directly be understood as translating a given amount of public expenditures (on, say, labour and capital) into a maximum possible amount of service poten-

⁴Alternative approaches to deal with this problem, often based on advanced econometric techniques, are reviewed and discussed with respect to productive efficiency in the education sector by Cordero-Ferrera et al. (2008).

tial (whereas the second stage is probably more appropriately analyzed in a supply-demand framework). This avoids the bias induced by using final outputs that are influenced by citizen coproduction (see above).⁵ Moreover, and importantly, focus on this first stage also mitigates identification concerns raised, among others, by Hoxby (1999, 2000). Indeed, analysing the effect of Tiebout choice on local education outcomes, she relies on an instrumental variables approach to separate variation "*driven by exogenous factors that affect the **supply** of school districts*" from that which "*is endogenous to observed student achievement or that is driven by the **demand** for school districts*" (Hoxby, 2000, 1210, boldface added). Our exclusion of demand-side factors by focusing on the first-stage of the public good production process in principle has an equivalent effect on the model's identification possibilities. Nonetheless, in the absence of an instrumental variables technique (which, in effect, has not been developed yet for the empirical model we introduce below), it should be clear that it only mitigates identification concerns and will not be able to fully resolve them.

Two potential limitations of our approach should be pointed out here. First, one might argue that thus far we implicitly assume public service providers to provide services fitting to local preferences. In other words, decisions regarding service potential are assumed to reflect the characteristics of expected demanders. For example, service providers are less likely to provide access during weekends in areas where everybody works weekends, while libraries probably buy more (less) children's books in communities with high birthrates (large share of elderly). This, however, need not hold in reality and we fully agree that a library providing unwanted services (e.g., books no one has an interest in reading, but which are the most inexpensive available) should not be designated as fulfilling its task, *even when* it provides the maximum possible amount of services for a given budget. Indeed, a more appropriate overall benchmark would be to consider how the service provider performs in terms of the production of services relevant for the socio-demographic make-up of the jurisdiction. However, this argument introduces the appropriateness of services or the responsiveness of the local service provider into the analysis (see also footnote 1), and thus goes beyond productive efficiency in a strict sense. Once again, it is important to stress that we are *not* claiming that these additional elements are less important than productive efficiency, but simply that analyses of pure productive efficiency should regard the appropriate framework and not *implicitly* encompass such effects.

Second, the clear distinction made above may not be all that clear in reality and public service providers could be argued to generally have some influence or control over the second stage of the production process. For example, librarians have an important role to facilitate

⁵Referring once more to Becker (1965) and Lancaster (1966), our focus on stage one of the public good production process is similar to arguing that the productive efficiency of private-sector firms is best evaluated using 'goods' and not 'commodities' as the firms' output. This is, unsurprisingly and uncontroversially, exactly what scholars of private-sector productive efficiency have done for decades.

or actively promote reading through the amount and quality of assistance programs, courses and/or lectures offered; fire services are involved in raising awareness of fire risk and prevention; and so on. Still, such demand-influencing activities by service providers could within our framework be seen as first-stage outputs (i.e. service potential). While they thus do not invalidate our suggestion to focus on stage one only (because accountability is much more clear-cut here than in stage two), they do potentially affect the interpretation of the inefficiency measured at stage one. Indeed, if this second-stage influence of the service provider is ignored or no appropriate measures for such often intangible services are included, the basic inputs used in the production of programmatic inputs are overstated and the resulting efficiency measure at stage one overestimates true inefficiency. The key requirement in avoiding this bias would be to identify those inputs specific to stage one alone, *or* include variables accounting for both tangible and intangible service potential in the analysis.⁶

2.2 Determinants of efficient public good provision: Hypotheses

Now that we have a clear understanding of the appropriate inputs and outputs when analyzing efficient provision of public goods, it remains to describe what might explain variations in efficiency across jurisdictions. Based on extent political economy and public choice literatures, we thereby concentrate on the institutional (i.e., political, socio-demographic as well as financial) environment in which public service providers (in this case, public libraries) operate. In the remainder of this section, we informally motivate the empirical application below by discussing how various local circumstances might be linked to local public library efficiency.

Firstly, the *political environment* is likely to matter because (1) right-wing parties and (2) male politicians might have different priorities compared to, respectively, left-wing governments and female politicians. This prediction follows from a large political economy literature stating that ideology determines politicians' policy preferences (e.g., Hibbs, 1977), as well as from more recent evidence indicating that gender is an important indicator of policy preferences (e.g., Lott and Kenny, 1999; Edlund and Pande, 2002) and determines policy outcomes (e.g., Chattopadhyay and Duflo, 2004; Geys and Revelli, 2009; Svaleryd, 2009). Specifically, this literature suggests that the latter groups - i.e. left-wing and female politicians - tend to be more egalitarian and socially conscious. Translated to the present setting, this could lead left-wing and female politicians to be more inclined to focus on non-economic benefits of libraries (i.e., supplying the opportunity to read for the less well-off, stimulating the population to read more extensively, etc.), thereby putting more pressure on local public libraries to maximize their service potential (given the limited budget). In other words, for a *given*

⁶Note that ignoring potential variation in the quality of service potential across service providers would have similar effects.

budget, they might be more likely to push for higher service potential.

Still, a second channel through which the political environment might matter is that left-wing parties often have close links to (public sector) trade unions. As a result, they might experience a stronger motivation to create 'rents' for public sector unions (Grossman and Helpman, 2001). This could result in shifting existing inputs towards labour (i.e. more staff, less books) or to increasing the library's overall budget to support additional employment (Mueller and Murrell, 1986). In either case, however, it is not immediately clear how such budget re-allocation or expansion would affect libraries' productive efficiency. Indeed, neither need imply a less productive efficient use of funds – unless one is willing to assume, say, that higher labour usage is necessarily detrimental to productive efficiency.

This discussion yields a first hypothesis:

- H1: Library efficiency is affected by a) the ideological stance of the local government and
b) the share of women in the local council.

The municipal *socio-demographic* make-up is likely to affect library efficiency through its influence on the demand and willingness to pay for cultural goods. Both of these generally rise with income (e.g., Throsby, 1994; Schulze and Ursprung, 1998), such that high-income residents “*may constitute a special interest group striving for cultural provisions*” (Werck *et al.*, 2008, 47). They may therefore pressure local public libraries to maximize their service potential given budgetary constraints. Related, distance is often argued to play a crucial role in deciding whether or not to visit a cultural event (e.g., Verhoeff, 1992; Boter *et al.*, 2005; De Graaff *et al.*, 2009; Bhatt, 2010). High concentration of population, by reducing the average cost of travelling to the library, thus increases the group of potential users of the library's services. Moreover, lower travel costs can be argued to increase the 'option value' of library services.⁷ As such, for a given population size, urban areas (which are more densely populated) may have a larger share of its population interested in and striving for the efficient public provision of cultural goods (in this case, library services). In both cases, one could interpret the variables' effects also in terms of a monitoring argument (see also below): i.e. with more people likely to use libraries, monitoring of how the library is using its resources is likely to increase.

We should note, however, that the possibility of congestion (for early analyses of congestion in public good provision, see, e.g., Brueckner, 1981; Hochman, 1982; Oates, 1988) may not only limit service provision when demand is very high (affecting the library's final output), but might also limit the force of the previous arguments. Indeed, expectation of congestion may well reduce individuals' propensity to pressure local service providers for

⁷This follows from standard option pricing theory. Specifically, a decrease in the 'strike price' of an option (e.g., through lower travel costs) to acquire a given underlying commodity (i.e., library services) increases its value (see also Werck *et al.*, 2008).

efficient provision by diminishing their expected benefit of the service: i.e. the benefit of the service becomes discounted by the probability of not achieving it due to congestion. This is supported by studies showing reduced willingness-to-pay as well as willingness-to-travel when expected congestion increases (e.g., McConnell, 1977; Menz and Mullen, 1981; Ashworth and Johnson, 1996; for a review relating to cultural resources, see Noonan, 2003).

This discussion leads to our second hypothesis:

H2: Library efficiency is affected by a jurisdictions' a) wealth and b) population density.

Finally, we know from principal-agent theory that monitoring reduces information asymmetries between principal and agent. Indeed, this literature shows that under perfect monitoring a first-best solution can still be reached. In other words, there is a strong disciplining effect of monitoring that - by limiting possibilities for wasteful spending and rent extraction by the agent - is strictly welfare-improving (Alchian and Demsetz, 1972; Jensen and Meckling, 1976; Holmström, 1979). Recent experimental evidence generally supports this prediction (e.g., Nagin *et al.*, 2002; Dickinson and Villeval, 2008). Importantly, monitoring is likely to be higher when a larger share of library spending derives from local sources (i.e., fees and municipal subsidies) rather than subsidies from higher-level governments. The reason lies in a form of fiscal illusion, whereby voters fail to fully understand that grants from higher-level governments have to be financed through tax revenues by these governments as well (and that they will at least partly provide these resources). Such an imperfect mapping of consumers and financiers of library services (or, in other words, when fiscal institutions are not built on the principle of 'fiscal equivalence'; Olson, 1969) reduces the incentive to act as efficiency guards. It is, after all, perceived to be other people's money that is being wasted, and voters may well care about government (in)efficiency only when they are directly confronted with the tax bill for public good provision. In other words, local public libraries' accountability to the public (for fees and charges) and higher-level governments (who provide grants) may differ. Applying this line of argument to our setting, libraries are "*more likely to value the careful use of public money when it originates mainly from own revenue sources rather than external transfers*" (Geys *et al.*, 2010, 266). This gives our third and final hypothesis (for similar arguments, see Hoxby, 2000; Hemmeter, 2006):

H3: Library efficiency is higher when resources derive to a larger extent from own revenues.

3 Empirical methodology

To estimate efficiency in the first stage of the library production process and to determine its politico-economy correlates (which are in the remainder considered as exogenous from the library's point of view), we could in principle employ several modelling techniques. However,

a closer look at the data and hypotheses limits the possibilities. First, we should focus on a non-parametric model as there is no *a priori* information on the appropriate production technology for public services. In other words, we have no reason to believe that the relationship between the inputs, outputs and exogenous characteristics follows a specific functional form (e.g., Cobb-Douglas, Translog, Fourier, ...). Although non-parametric models have a lower rate of convergence, they have been shown to be more consistent compared to wrongly specified parametric models (Kneip *et al.*, 1998). Second, as we have no information on price variables, we have to rely on a branch of non-parametric models particularly designed for public performance analysis: Data Envelopment Analysis (DEA; Charnes *et al.*, 1978) and Free Disposal Hull (FDH; Deprins *et al.*, 1984). Finally, given that we want to control for heterogeneity and test for the influence of the exogenous environment on productive efficiency, the choice of modelling techniques is further narrowed to conditional efficiency estimators (Daraio and Simar, 2005, 2007) and, in particular, conditional efficiency models that allow for discrete and continuous exogenous variables (De Witte and Kortelainen, 2008).

The model starts from the set χ_n of observed combinations of inputs x ($x \in \mathbb{R}_+^p$) and outputs y ($y \in \mathbb{R}_+^q$). The set of all feasible input-output combinations defines the production technology: $\Psi = \{(x, y) \in \mathbb{R}_+^{p+q} \mid x \text{ can produce } y\}$. To determine the efficiency of the evaluated libraries, we start from the best practice observations, i.e., the libraries that are using the least inputs x for a given amount of outputs y (this is the so-called input-orientation; for alternative orientations, see Fried *et al.*, 2008). These best practice observations constitute the best practice frontier, i.e., the border of the production technology Ψ . The inefficiency θ of the evaluated entity (x, y) is estimated as the distance to the best practice frontier:

$$\theta(x, y) = \inf \{\theta \mid (\theta x, y) \in \Psi\} \quad (1)$$

where the input efficiency measure $\theta(x, y) \leq 1$ is the proportionate decrease of inputs, which the library operating at level (x, y) should attain in order to be considered ‘efficient’ (i.e., $\theta(x, y) = 1$).

Two options now arise. One could impose convexity on the production possibilities (as in DEA) or not (as in FDH). Not imposing convexity clearly implies a more general approach. Moreover, there are “*no valid theoretical arguments for assuming a priori that production possibilities are truly convex*” (Cherchye *et al.*, 2000, 263-264) and some empirical studies suggest violations of the convexity hypothesis (e.g., Hasenkamp, 1976). Hence, as there is no clear justification in our application to estimate a convex hull around the data, we concentrate on the FDH model. The FDH model estimates the production possibility set as:

$$\hat{\Psi}_{FDH} = \{(x, y) \in \mathbb{R}_+^{p+q} \mid y \leq y_i, x \geq x_i, (x_i, y_i) \in \chi_n\}. \quad (2)$$

The FDH estimator for the Farrell input-oriented efficiency score is obtained by replacing Ψ with $\hat{\Psi}$ in equation (1).

However, a major disadvantage of the traditional non-parametric FDH model is that all n observations in the sample χ_n are considered to be potential best practices: $Prob((x, y) \in \Psi) = 1$. Therefore, atypical observations (e.g., due to measurement errors, very atypical structure of the entity, and so on) heavily influence the best practice frontier and, as a direct result, the efficiency scores. To reduce the influence of these atypical observations, we follow Cazals *et al.* (2002) in estimating the FDH efficiency of equation (1) relative to a partial frontier constituting of $m < n$ observations. By repeatedly drawing (B times) with replacement a subset of m observations among those x_i such that $y_i \geq y$ and averaging the efficiency scores relative to these B subsets, we obtain a so-called robust efficiency estimate $\theta^m(x, y)$ [robust in the sense that the efficiency scores are more robust to outlying observations]. The robust estimates $\theta^m(x, y)$ are no longer bounded by 1 as the evaluated observation is not always included in the reference set. These ‘super-efficient’ efficiency scores (i.e., if $\theta^m(x, y) < 1$) indicate that the observation is using less inputs than the average m evaluated observations in its reference set. As such, the super-efficient observation is doing better than what would be expected.⁸

The robust efficiency approach of Cazals *et al.* (2002) proves extremely convenient to incorporate the exogenous environment. Traditional non-parametric models suffer from a separability condition in that the operational environment is assumed not to influence the inputs and outputs. However, in real life applications, this is clearly unrealistic. The conditional efficiency approach, developed by Cazals *et al.* (2002) and Daraio and Simar (2005, 2007), allows to incorporate the exogenous environment. Basically, while using the robust efficiency model described above, the idea is to draw the subsample of size m in such a way that similar observations have a higher probability of being drawn. In practice, one obtains weights by estimating a kernel density around the evaluated exogenous characteristics. Hence, the resulting ‘conditional’ efficiency estimates $\theta^m(x, y | z)$ compare like with likes. As the seminal contributions did not allow for multivariate analysis of both discrete and continuous exogenous variables, De Witte and Kortelainen (2008) extended the approach. Basically, their approach uses mixed (i.e., both discrete and continuous) Kernel smoothing around the exogenous variables such that for every observation the probability of being similar to the evaluated observation is known.

A second advantage of the De Witte and Kortelainen (2008) extension - crucial for our analysis - arises from the possibility for statistical inference concerning the exogenous variables. This extends the original contribution of Daraio and Simar (2005), which allowed for a graphical inference (i.e., favorable or unfavorable), to estimating a non-parametric p -value (using an on bootstrapping based approach). Similar as in Daraio and Simar (2005, 2007),

⁸Following the literature (e.g., Daraio and Simar, 2007), we select the size of the partial frontier m as the value of m as of which the percentage of super-efficient observations is only decreasing marginally with m . In our application, this corresponds to $m = 50$ (whereas $n = 290$, see below).

the procedure is implemented by estimating the relation between the exogenous variables and the ratio of the conditional $\theta^m(x, y | z)$ to the unconditional $\theta^m(x, y)$ efficiency. Indeed, if an exogenous variable shows an unfavorable correlation with performance, then $\theta^m(x, y | z)$ (i.e., efficiency when taking z into account) will be larger than the unconditional efficiency $\theta^m(x, y)$ for large values of z compared to small values of z (Daraio and Simar, 2007). Non-parametrically bootstrapping this non-parametric regression allows us to obtain statistical inference (in particular, p -values) on the correlation between the environmental variable and the efficiency score (which, clearly, does not necessarily constitute a causal effect).⁹

4 Institutional setting and data

Our empirical application exploits an exceptionally rich dataset of (all 290) municipal public libraries in Flanders in 2007. The data derive from the Department ‘Social Development and Local Cultural Policy’ (Afdeling Volksontwikkeling en Lokaal Cultuurbeleid) of the Flemish Regional government. They collect - and make publicly available - information on library revenues (e.g., subsidies, fines and fees), expenditures (on personnel, infrastructure, library collection maintenance), collection size (e.g., books, CDs, DVDs, and so on) and operations (i.e., circulation, requests, and so on) since 1998. We employ the most recent data available (i.e., 2007) as a change in the data collection methodology in 2006 makes the resulting data imperfectly comparable across time. Data on opening hours are unfortunately not centrally collected and have been brought together by contacting all 290 libraries.

Given that FDH-based approaches - as the one employed here - tend to be sensitive to the number of inputs and outputs included (inclusion of more inputs and outputs increases the number of efficient observations; see Kneip *et al.*, 1998), we opt for three input and four output variables. As inputs, we use expenditures on (1) personnel, (2) operating expenditures (Opex; mainly maintenance of the collection) and (3) infrastructure.¹⁰ It is important to note at this point that charitable donations to and employment of volunteers in public libraries are uncommon in Flanders (in contrast to, for example, the US) and are, as such, not included in the analysis. These inputs, which fully exhaust the library expenditure budget, are used to provide (1) youth books, (2) fiction and non-fiction books and (3) other media (CD, DVD, VHS, CD-ROM) during (4) a given number of hours per week. Hence, we use three

⁹Causality and endogeneity issues are traditionally dealt with by a parametric instrumental variables (IV) approach. However, in a non-parametric setting, IV are technically impossible as efficiency estimators allowing for instruments are not yet developed (neither for the non-parametric estimator used here, as for the parametric / semi-parametric alternatives available in the literature; see Fried *et al.*, 2008).

¹⁰This infrastructure spending does *not* refer to big investment projects (such as major renovations or additions to the library buildings), which tend to be lumpy and time-specific. Instead, it measures the annual, contemporaneous expenditure on infrastructure that occurs because books must be housed in an enclosed space and larger book collections require a larger space with higher maintenance costs.

collection-related variables (expressed in number of books) as indicators of library service potential and add, as a fourth output, the total number of opening hours per week (since this proxies the actual accessibility of the library collection for potential borrowers).¹¹ We should thereby note that although the three former output variables correspond to stock measures, which may benefit older libraries, the variable returns to scale approach employed in our FDH model smoothly accounts for this (see Fried *et al.*, 2008). Descriptive statistics of the variables employed are presented in Table 1.

Table 1: Descriptive statistics for 2007 (n=290)

	Average	St. Dev	Min	Max
Inputs (in €)				
Personnel	366539	474973	52914	4698859
Opex	70725	80184	5897	609432
Infrastructure	66827	217529	103	1794090
Outputs (in absolute amounts)				
Opening hours per week	24	10	10	51
Youth books	25120	20800	4602	161986
Fiction and non-fiction books	40852	38909	6806	285218
Media (CD, DVD, VHS, CD-ROM)	33455	59187	0	523144
Operational environment				
Ideological complexion ICG (0-10 scale)	5.022	0.727	2.500	6.300
Female in council (%)	0.336	0.078	0.080	0.600
Female mayor (dummy)	0.093	0.291	0	1
Income (in 1000€)	12.930	1.390	9.547	17.536
Population (total in absolute amount)	19780	20034	2337	235143
Population density (per km ²)	527	427	62	3053
Subsidies Flanders (%)	0,833	0,102	0,165	0,979

To evaluate the three central hypotheses derived in Section 3.2, we examine the (1) ideological stance of the local government, (2) share of women in the local council, (3) wealth of the municipality, (4) population concentration, and (5) source of public funding. These are measured as follows:

First, we measure the Ideological Complexion of the local Government (ICG) as $ICG =$

¹¹Clearly, the service potential of a library goes beyond these four variables and can be thought to also include the amount and quality of assistance programs, courses, lectures and/or exhibitions offered. Unfortunately, however, data for such outputs are unavailable. As such, to the extent that basic inputs are employed for the provision of such services, our analysis is likely to over-estimate true technical inefficiency in stage one of the public service production process (see above).

$\sum_{i=1}^n (p_i \cdot \text{Complexion}_i)$, where p_i is the seat share of party i in the College of Mayor and Aldermen (i.e., the local government) and Complexion_i refers to the ideological position of this party on a Left-Right scale (from 0 to 10). The data concerning a party’s ideological position were obtained from Buelens *et al.* (2008) and are based on a self-placement survey asking presidents and spokesmen of the parties in the municipalities to locate their party on an ideological scale between 0 (Left) and 10 (Right). The figures range from 2.0 (Groen!) to 6.3 (Open VLD) (the extreme-right-wing party Vlaams Belang was not represented in any local government and is therefore not in the dataset).

Second, to measure the influence of female representation, we use two operationalizations. The first measures female representation as the share of female members in the municipal council (the local parliament). The more women are elected into the council, the more likely it is that female preferences are translated into actual policies. The second operationalization attempts to establish whether female preferences are more likely to come about when there is a female mayor. In this case, we define female representation using an indicator variable equal to 1 when the mayor is female, 0 otherwise.¹²

Third, real taxable per capita income (in €1000) is included to assess whether efficiency differs in wealthier municipalities. Population density (measured as inhabitants per km²) is taken up as a measure for the degree of urbanization.

Finally, libraries are a heavily subsidized public service in Flanders. A large share of these subsidies derives from the Flemish Regional government, which provides each municipality with a ‘basic subsidy’ of 6€ per inhabitant (with a minimum of 50,000€) towards the financing of its library personnel. These subsidies constitute no less than 83% of total library revenues on average (though varying between 16.5% and 97.9%; see Table 1). The remaining revenue is provided by municipal subsidies and various alternative resources such as membership fees, borrowing fees and fines. We include the share of regional subsidies in the total library budget to evaluate the third hypothesis, i.e. whether efficiency is higher (lower) when resources derive to a larger extent from own revenues (higher-level government subsidies).

5 Results

The results of the efficiency estimations are summarized in Table 2. In column 1, we present the results when we do not account for the exogenous environment (i.e., ‘unconditional’ efficiency). We find an average efficiency score of 0.79. This indicates that, on average, Flemish local public libraries would have to decrease their inputs by approximately 21% in

¹²Still, the data do not allow us to interpret our findings as causal evidence in favour of female influence. Indeed, women might simply get voted into office more often in municipalities that have a more ‘egalitarian’ population. As such, it might be underlying differences in popular preferences (rather than women’s policy preferences as such) causing both a higher share of female councillors and more efficient libraries.

order to produce their outputs equally efficient as the best practices. There is, however, a very large variation in the performance across libraries, as can be seen from the sizeable standard deviation around this average inefficiency. Moreover, some efficiency scores are significantly larger than 1 (i.e., $\theta^m(x, y) < 1$). Hence, some observations can be viewed as super-efficient: they perform better than the average m observations in their reference sample.

To examine Hypotheses 1 to 3, we develop five alternative conditional efficiency models. In Model 1, we examine the correlation between the efficiency score and ideological stance, share of female politicians in the local council, average income, and population concentration via a nonparametric bootstrap procedure (see above). Model 2 adds population size to this baseline model, in order to check whether any population concentration effect not merely derives from a larger population as such. Model 3 adds a dummy variable for female mayors to assess whether this adds to the explanatory power of the model after controlling for overall female representation in the local council (as in Model 1). To test Hypothesis 3, we include the percentage of regional subsidies in total library revenues, while controlling for the ideological preferences (in Model 4) and average income (Model 5) of the municipality. A systematic presentation is given in Table 3.

Once we account for the exogenous environment (i.e., conditional efficiency), our earlier conclusions change in two important ways (results summarized in columns 2 to 6 of Table 2). First, the average efficiency score no longer significantly deviates from 1. Second, the standard deviation around this mean reduces significantly. Both results indicate that a large part of the variation in inefficiency observed in the unconditional efficiency estimates is related to differences in the exogenous factors introduced.

These summarized results, however, do not allow any statements regarding Hypothesis 1-3. Thus, we represent the full estimation results for all five conditional models in Table 3. We thereby consider the effect on the median, rather than the mean, as the former is less influenced by extreme values. Due to the structure of the non-parametric bootstrap, we only present whether the exogenous variable is significantly (un)favorably correlated with efficiency (since the marginal coefficient on the median is less meaningful; see De Witte and Kortelainen, 2008). Finally, before discussing the findings, we should repeat that care should be taken with a causal interpretation of our findings. While our specific case selection and the separating out of demand-side effects mitigates concerns about endogeneity and identification (see above), the impossibility of employing an IV approach (see footnote 9 above) should induce due caution.

A first finding observed throughout all specifications Table 3 is that right-wing councils are associated with a statistically significantly lower level of public library productive efficiency, *ceteris paribus*. Such finding is consistent with the idea that left-wing governments care more about maximizing library service potential (under a given budget). It appears at odds with

Table 2: Order-m efficiency score (N=290)

	Unconditional (robust FDH)	Conditional				
		Model 1	Model 2	Model 3	Model 4	Model 5
Average	0.7892	1.0507	0.9891	1.0078	1.1082	1.0096
St. Dev.	0.6575	0.1946	0.1294	0.1606	0.2394	0.2029
Min	0.1965	0.5315	0.4778	0.5061	0.5018	0.5019
Max	5.4843	2.4027	2.0493	2.4284	2.0760	2.0093

Table 3: Influence of the exogenous environment on service potential of libraries

	Model 1	Model 2	Model 3	Model 4	Model 5
ICG	unfavorable (<2E-16) ***	unfavorable (<2E-16) ***	unfavorable (<2E-16) ***	unfavorable (0.0783 *	unfavorable (0.060) *
Female council	favorable (0.235)	favorable (0.372)	unfavorable (0.540)		
Income	favorable (0.025) **	favorable (<2E-16) ***	favorable (0.020) **		favorable (<2 E-6) ***
Population density	favorable (0.090) *	favorable (<2E-16) ***	favorable (0.010) ***		
Population (total)		favorable (0.830)			
Female mayor			unfavorable (0.170)		
Regional subsidies				unfavorable (0.045) **	unfavorable (0.057) *

Note: n=290; Bootstrapped p-values between brackets; ***, ** and * denote significance at 1%, 5% and 10%-level.

a prediction under which such governments would be mainly interested in expanding local public budgets to appease public-sector unions (although such hypothesis would also require the assumption that unions are rent-seekers, which, obviously, need not be true).

Secondly, our findings are also consistent with the relation brought forward in Hypothesis 2. Municipalities with higher income and urbanization levels (i.e., more densely populated) have better performing libraries. The latter is not driven by population size. Indeed, when including population size in the estimation (Model 2), this turns out to be insignificant, while the statistical significance of population concentration remains unaffected.

Thirdly, our results for the regional subsidies variable are in line with the expectation expressed in Hypothesis 3. Libraries turn out to have a higher service potential for a given budget in municipalities where library revenues derive to a larger extent from local resources (i.e., municipal subsidies, fees and fines). This could indicate that incentives for monitoring are undermined when a larger share of income derives from higher-level government subsidies. This supports Hemmeter's (2006) findings for a sample of 3308 US library systems. It also links to recent findings by Geys *et al.* (2010, 265), who show that "*the efficiency-enhancing effect of voter involvement*" in German municipalities is significantly reduced when a larger share of local revenues comes from intergovernmental subsidies.

Finally, and somewhat at odds with a recent and quickly expanding literature linking female representation to changes in public policy outcomes, we cannot establish a robust nor significant link between the share of female representatives in the local council and library efficiency. That is, although the share of female representatives is positively related to median efficiency in two of the three models where it is included, its coefficient estimate always fails to be statistically significant. A similar absence of statistical significance is reached when examining the gender of the mayor (Model 3). Female mayors are not associated with a significantly higher level of library performance, *ceteris paribus*.¹³

6 Conclusion

Economic efficiency – in terms of maximizing output for a given level of inputs (e.g., Koopmans, 1951; Fried *et al.*, 2008) – has recently become an increasingly important element in public good provision. This has led to a concomitant increase in scholarly attention for the determinants of such efficiency. We added to this burgeoning literature in three ways.

First, we characterized public good provision as a two-stage production process. One stage translates basic inputs into service potential, while a second stage describes how these

¹³One potential explanation for this non-finding might be that municipalities with high female representation also tend to have more left-wing governments (assuming women are more likely to run - and be elected - on left-wing party lists). This, however, is not supported by the data. Indeed, while we do find the expected negative correlation between ICG and fcouncil, this relation is rather weak ($r=-0.1140$; $p=0.07$).

are transformed into observed outcomes. This separation is crucial to accurately defining inputs and outputs in the analysis of productive efficiency (and thus determines the accuracy of the results from such analyses). Indeed, as the second stage outcomes are influenced by public demand (and thus in part lie beyond the control of the public service provider), they are inappropriate to evaluate the pure productive efficiency of public good provision. Hence, we argued that one should concentrate on the first stage (in which the production of service potential is fully under control of the service provider).

Second, we employed a specially tailored and fully non-parametric framework, which is rooted in popular Data Envelopment Analysis (DEA) models (Charnes *et al.*, 1978). By using a robust (i.e., allowing for outlying observations; Cazals *et al.*, 2002) and conditional (i.e., allowing for heterogeneity; Daraio and Simar, 2005, 2007) efficiency framework, we were able to evaluate non-parametrically (i.e., without any *a priori* assumption on the production function) how discrete and continuous exogenous variables are correlated with productive efficiency (although this does not necessarily allow causal inferences). We thereby included the operational environment immediately in the efficiency estimates, thus avoiding use of a separability condition, which inappropriately assumes that the exogenous environment is not linked with the inputs and outputs.

Third, relying on a large political economy and public choice literature, we used the above framework to examine how political economy factors are related to efficiency of local public good provision. Our findings - using data on 290 Flemish local public libraries - suggest that library productive efficiency is higher when (1) the ideological stance of the local government is more left-wing, (2) the population is wealthier, (3) the area is more densely populated and (4) public service revenues derive to a larger extent from local resources. In addition, our results indicate that, *ceteris paribus*, (5) population size, (6) the share of female representatives in the local council, and (7) having a female mayor shows no significant correlation to library performances.

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